



Banco de Portugal

EUROSISTEMA

Estudos e Documentos de Trabalho

Working Papers

13 | 2007

EXACT LIMIT OF THE EXPECTED PERIODOGRAM IN THE
UNIT-ROOT CASE

João Valle e Azevedo

September 2007

The analyses, opinions and findings of these papers represent the views of the authors, they are not necessarily those of the Banco de Portugal.

Please address correspondence to

João Valle e Azevedo

Economics and Research Department

Banco de Portugal, Av. Almirante Reis no. 71, 1150-012 Lisboa, Portugal;

Tel.: 351 21 3130163, Email: jvazevedo@bportugal.pt

BANCO DE PORTUGAL

Economics and Research Department

Av. Almirante Reis, 71-6th floor

1150-012 Lisboa

www.bportugal.pt

Printed and distributed by

Administrative Services Department

Av. Almirante Reis, 71-2nd floor

1150-012 Lisboa

Number of copies printed

200 issues

Legal Deposit no. 3664/83

ISSN 0870-0117

ISBN 978-989-8061-10-2

Exact Limit of the Expected Periodogram in the Unit-Root case

João Valle e Azevedo*

Banco de Portugal & Universidade NOVA de Lisboa

September 21, 2007

Abstract

We derive the limit of the expected periodogram in the unit-root case under general conditions. This function is seen to be independent of time, thus sharing a fundamental property with the stationary case equivalent. We discuss the consequences of this result to the frequency domain interpretation of filtered integrated time series.

JEL Classification: C22

Keywords: Periodogram, Unit root

1 Introduction

Solo (1992) has shown that certain continuous-time stationary increment processes possess many of the frequency domain properties of stationary processes. Crucially, although their variance is infinite or time-varying (depending on the specification of initial conditions), they have a time-invariant spectrum, defined there as the limit of the expected periodogram. This more general definition of spectrum helps us understand the frequency domain properties of certain non-stationary processes, circumventing the restrictive nature of the standard spectral representation theorems for stationary processes.

*Address: Av. Almirante Reis, 71-6th floor, 1150-012 Lisboa, Portugal; E-mail: jvazevedo@bportugal.pt; Phone: +351 213130163. Most research was done while I was a graduate student at Stanford University. I gratefully acknowledge the Portuguese Government, through the Fundação para a Ciência e Tecnologia, for financial support during my graduate studies.

We show in this paper that Solo's (1992) main result holds in the case of (discrete-time) time series processes containing one unit root. Under very general conditions, we provide exact expressions for the time-invariant spectrum of an integrated time series, defined as the limit of the expected periodogram. It is shown that this limit differs from the commonly defined (pseudo-) spectrum of an integrated time series. We will discuss the nuisance that this fact represents to the interpretation of the consequences of applying linear filters that render the series stationary.

2 The limit of the expected periodogram

Denote $I_{T,x}(\omega_j)$ as the periodogram of the sequence $\{x_t\}_{t=1}^T$, where $\omega_j = 2\pi j/T$ are the integer multiples of $2\pi/T$ that fall in the interval $] - \pi, \pi]$. Restricting ourselves to real sequences and noting that $I_{T,x}(\omega_j) = I_{T,x}(-\omega_j)$ in this case, we extend as usual the periodogram for every frequency in the interval $[-\pi, \pi]$ in the following way:

$$I_{T,x}(\omega) = \begin{cases} I_{T,x}(\omega_k), & \omega_k - \pi/T < \omega \leq \omega_k + \pi/T \\ I_{T,x}(-\omega), & 0 < \omega \leq \pi \end{cases}$$

For $\omega \in [0, \pi]$, let $g(T, \omega)$ be the multiple of $2\pi/T$ closest to ω . If $\omega \in [-\pi, 0[$, let $g(T, \omega) = g(T, -\omega)$. Then,

$$I_{T,x}(\omega) = I_{T,x}(g(T, \omega)) \tag{1}$$

If $\{x_t\}_{t=1}^T$ is a sample from a stationary time series with mean μ and the autocovariance function $\gamma(\cdot)$ is absolutely summable, it can be shown (see, e.g., Brockwell and Davis, 1991, p.343) that:

$$E[I_{T,x}(0) - T\mu^2] \rightarrow 2\pi S_x(0) \text{ as } T \rightarrow \infty \tag{2}$$

$$E[I_{T,x}(\omega)] \rightarrow 2\pi S_x(\omega) \text{ as } T \rightarrow \infty, \omega \neq 0$$

where $S_x(\omega)$ is the spectrum of x_t . That is, when the sample size grows the periodogram converges to the distribution of variance as revealed by the spectral representation theorem. As Solo (1992), in the analysis of the spectrum of continuous-time, stationary increments processes, we argue that the result in (2) is a less restrictive inversion relation than that implied by the spectral representation theorem. The question that we address is whether or not the relation in (2) remains valid in the case of integrated processes. Does the expected value of the periodogram of an integrated series, which can be seen as a distribution of power, converge to a time-invariant function? The surprising answer is that it does, at least if the order of integration is 1 and for a

very broad class of stationary increments. This is summarised in theorem 1.

Theorem 1. Let $u_t = \psi(L)\varepsilon_t = \sum_{j=-\infty}^{\infty} \psi_j \varepsilon_{t-j}$, where $\sum_{j=-\infty}^{\infty} |\psi_j| |j|^{\frac{1}{2}} < \infty$ and $\{\varepsilon_t\}$ is a white noise sequence such that $E[\varepsilon_t] = 0$ and $\text{Var}[\varepsilon_t] = \sigma_\varepsilon^2 < \infty$. Consider the process $\{x_t\}$ verifying $x_t - x_{t-1} = u_t, \forall t$. Then, the periodogram of x_t , $I_{T,x}(\omega)$, has the following properties:

$$\begin{aligned} i) \quad & T^{-2}E[I_{T,x}(0)] \rightarrow \frac{2\pi}{3}S_{\Delta x}(0) \text{ as } T \rightarrow \infty, \text{ assuming } x_0 = 0 \\ ii) \quad & E[I_{T,x}(\omega)] \rightarrow 2\pi S_x(\omega) \text{ as } T \rightarrow \infty, \omega \neq 0 \end{aligned} \quad (3)$$

where $S_{\Delta x}(0)$ is the spectrum of $x_t - x_{t-1} = u_t$ at zero frequency and

$$S_x(\omega) = \frac{\sigma_\varepsilon^2 (|\psi(e^{-i\omega})|^2 + |\psi(1)|^2)}{2\pi |1 - e^{-i\omega}|^2}, \omega \neq 0 \quad (4)$$

Proof: Consider first the case $\omega \neq 0$. Fix any $\omega \in]0, \pi[$. Then by (1) $I_{T,x}(\omega) = I_{T,x}(\omega_j)$ for some Fourier frequency ω_j . The discrete Fourier transform of $x_t - x_{t-1} = \Delta x_t$, denoted by $J_{T,\Delta x}(\omega_j)$, can be decomposed in the following way:

$$\begin{aligned} J_{T,\Delta x}(\omega_j) &= T^{-\frac{1}{2}} \sum_{t=1}^T \Delta x_t e^{-i\omega_j t} = T^{-\frac{1}{2}} \sum_{t=1}^T (x_t - x_{t-1}) e^{-i\omega_j t} = \\ &= T^{-\frac{1}{2}} (1 - e^{-i\omega_j}) \sum_{t=1}^T x_t e^{-i\omega_j t} + T^{-\frac{1}{2}} \sum_{t=1}^T x_t e^{-i\omega_j(t+1)} - T^{-\frac{1}{2}} \sum_{t=1}^T x_{t-1} e^{-i\omega_j t} = \\ &= T^{-\frac{1}{2}} (1 - e^{-i\omega_j}) \sum_{t=1}^T x_t e^{-i\omega_j t} + T^{-\frac{1}{2}} (x_T e^{-i\omega_j(T+1)} - x_0 e^{-i\omega_j}) = \\ &= (1 - e^{-i\omega_j}) J_{T,x}(\omega_j) + T^{-\frac{1}{2}} (x_T e^{-i\omega_j(T+1)} - x_0 e^{-i\omega_j}). \end{aligned} \quad (5)$$

where $J_{T,x}(\omega_j)$ denotes the discrete Fourier transform of x_t . Now, the periodogram of Δx_t can be written as $I_{T,\Delta x}(\omega_j) = J_{T,\Delta x}(\omega_j) J_{T,\Delta x}(-\omega_j)$. Multiplying both sides of (5) by $J_{T,\Delta x}(-\omega_j)$, using the fact that $e^{-i\omega_j(T+1)} = e^{-i\omega_j}$ for the Fourier frequencies ω_j and rearranging terms we get:

$$\begin{aligned} |1 - e^{-i\omega_j}|^2 I_{T,x}(\omega_j) &= I_{T,\Delta x}(\omega_j) + T^{-1} (x_T - x_0)^2 - \\ &- J_{T,\Delta x}(\omega_j) T^{-\frac{1}{2}} (x_T - x_0) e^{-i\omega_j} - J_{T,\Delta x}(-\omega_j) T^{-\frac{1}{2}} (x_T - x_0) e^{i\omega_j} \end{aligned} \quad (6)$$

Now put $R_T(\omega_j) = J_{T,\Delta x}(\omega_j)T^{-\frac{1}{2}}(x_T - x_0)e^{-i\omega_j}$. Taking expectations we get:

$$E[R_T(\omega_j)] = T^{-1}e^{-i\omega_j}E\left[\left(\sum_{t=1}^T u_t e^{-i\omega_j t}\right)\sum_{t=1}^T u_t\right] = T^{-1}e^{-i\omega_j}\mathbf{1}'E[\mathbf{u}\mathbf{u}']\mathbf{e}$$

where $\mathbf{1}$ is a vector of ones, $\mathbf{e} = (e^{-i\omega_j}, e^{-2i\omega_j}, \dots, e^{-Ti\omega_j})'$ and $\mathbf{u} = (u_1, u_2, \dots, u_T)'$. Since $E[\mathbf{u}\mathbf{u}'] = [\gamma_{\Delta x}(j-i)]_{i,j=1}^T$, where $\gamma_{\Delta x}(\cdot)$ is the autocovariance function of Δx_t , we get finally:

$$E[R_T(\omega_j)] = T^{-1}e^{-i\omega_j}\sum_{l=1}^T\sum_{h=1}^T\gamma_{\Delta x}(h-l)e^{-i\omega_j h}$$

This can be decomposed as follows:

$$E[R_T(\omega_j)] = T^{-1}e^{-i\omega_j}\left(\sum_{h=0}^{T-1}\gamma_{\Delta x}(h)e^{-i\omega_j h}\sum_{l=1}^{T-h}e^{-i\omega_j l} + \sum_{h=-T+1}^{-1}\gamma_{\Delta x}(h)e^{-i\omega_j h}\sum_{l=1-h}^T e^{-i\omega_j l}\right) \quad (7)$$

Now, for $0 \leq h \leq T-1$ we have:

$$\left|\sum_{l=1}^{T-h}e^{-i\omega_j l}\right| = \left|\sum_{l=1}^T e^{-i\omega_j l} - \sum_{l=T-h+1}^T e^{-i\omega_j l}\right| = \left|0 - \sum_{l=T-h+1}^T e^{-i\omega_j l}\right| \leq h$$

since $\sum_{h=1}^T e^{-i\omega_j h} = \frac{1-e^{-i\omega_j T}}{1-e^{-i\omega_j}}e^{-i\omega_j} = 0$, as $e^{-i\omega_j T} = 1$ for the Fourier frequencies ω_j . The inequality follows from the fact that $\omega_j \neq 0$ or 2π . Also, for $-T+1 \leq h \leq -1$ we can conclude that:

$$\left|\sum_{l=1-h}^T e^{-i\omega_j l}\right| \leq |h|$$

All this means that we can bound (7) by

$$\begin{aligned} T^{-1}\sum_{|h|<T}|\gamma_{\Delta x}(h)||h| &\leq T^{-1}\sum_{|h|<T}\sum_{j=-\infty}^{\infty}|\psi_j\psi_{j+h}||h| \\ &\leq T^{-\frac{1}{2}}\sum_{|h|<T}\sum_{j=-\infty}^{\infty}|\psi_j\psi_{j+h}||h|^{\frac{1}{2}} \leq T^{-\frac{1}{2}}\left(\sum_{h=-\infty}^{\infty}\sum_{j=-\infty}^{\infty}|\psi_j\psi_{j+h}||h+j|^{\frac{1}{2}} + \sum_{h=-\infty}^{\infty}\sum_{j=-\infty}^{\infty}|\psi_j\psi_{j+h}||j|^{\frac{1}{2}}\right) \end{aligned}$$

$$= 2T^{-\frac{1}{2}} \left(\sum_{h=-\infty}^{\infty} |\psi_h| |h|^{\frac{1}{2}} \right) \left(\sum_{j=-\infty}^{\infty} |\psi_j| \right) \rightarrow 0 \text{ as } T \rightarrow \infty$$

since both series are convergent. Performing the same exercise for $R_T(-\omega_j)$ we conclude that the expected value of the last two terms in (6) converges to 0:

$$E[R_T(\omega_j) + R_T(-\omega_j)] \rightarrow 0 \text{ as } T \rightarrow \infty \quad (8)$$

As for the second term in (6) we get:

$$\begin{aligned} E[T^{-1}(x_T - x_0)^2] &= T^{-1} \text{Var} \left[\sum_{t=1}^T \Delta x_t \right] = \\ &= \sum_{|k| < T} \left(1 - \frac{|k|}{T} \right) \gamma_{\Delta x}(k) \rightarrow \sum_{k=-\infty}^{\infty} \gamma_{\Delta x}(k) = 2\pi S_{\Delta x}(0) \text{ as } T \rightarrow \infty \end{aligned} \quad (9)$$

by the dominated convergence theorem. Using (8), (9), the fact that $I_{T,\Delta x}(\omega) \rightarrow 2\pi S_{\Delta x}(\omega)$ as $T \rightarrow \infty$ (see (2)) and finally the fact that $|1 - e^{-ig(T,\omega)}|^2 \rightarrow |1 - e^{-i\omega}|^2$ as $T \rightarrow \infty$ (since $g(T,\omega) \rightarrow \omega$) we conclude that:

$$E\left[\frac{1}{2\pi} I_{T,x}(\omega)\right] \rightarrow \frac{S_{\Delta x}(\omega) + S_{\Delta x}(0)}{|1 - e^{-i\omega}|^2} = \frac{\sigma_\varepsilon^2 (|\psi(e^{-i\omega})|^2 + |\psi(1)|^2)}{2\pi |1 - e^{-i\omega}|^2}, \text{ as } T \rightarrow \infty, \omega \neq 0$$

which is time-invariant! For $\omega = 0$, we need to normalise the periodogram by T^3 instead of T , and also to take into account the initial condition $x_0 = 0$, which is equivalent to analyse the periodogram for $\{x_t - x_0\}$ instead of $\{x_t\}$. We get:

$$T^{-2} E[I_{T,x}(0)] = E\left[T^{-3} \sum_{t=1}^T (x_t - x_0) \sum_{t=1}^T (x_t - x_0)\right] = E\left[T^{-3} \sum_{t=1}^T \sum_{l=1}^t u_l \sum_{t=1}^T \sum_{l=1}^t u_l\right] = \mathbf{1}' E[\mathbf{u}_c \mathbf{u}_c'] \mathbf{1}$$

where $\mathbf{1}$ is a vector of ones and $\mathbf{u}_c = (u_1, u_1 + u_2, \dots, \sum_{l=1}^T u_l)$. Evaluating $E[\mathbf{u}_c \mathbf{u}_c']$ we conclude that:

$$T^{-2} E[I_{T,x}(0)] = T^{-3} \sum_{|k| < T} \gamma_{\Delta x}(k) \sum_{h=1}^{T-|k|} (|k| + h)h$$

But:

$$\begin{aligned} \sum_{h=1}^{T-|k|} (|k| + h)h &= \sum_{h=1}^{T-|k|} h^2 + |k| \sum_{h=1}^{T-|k|} h = \\ &= \frac{(T - |k|)(T - |k| + 1)(2(T - |k|) + 1)}{6} + \frac{(T - |k|)(T - |k| + 1)|k|}{2} \end{aligned}$$

Thus,

$$T^{-2}E[I_{T,x}(0)] = \frac{1}{3} \sum_{|k| < T} \gamma_{\Delta x}(k)R(T, |k|)$$

where, for fixed $|k|$, $\lim_{T \rightarrow \infty} R(T, |k|) = 1$. From the dominated convergence theorem, we finally conclude:

$$T^{-2}E[I_{T,x}(0)] \rightarrow \frac{2\pi}{3}S_{\Delta x}(0) \text{ as } T \rightarrow \infty$$

■

Example: Random walk. If $\{x_t\}$ verifies $x_t - x_{t-1} = \varepsilon_t, \forall t$ where $\{\varepsilon_t\}$ is a white noise sequence such that $E[\varepsilon_t] = 0$ and $Var[\varepsilon_t] = \sigma_\varepsilon^2$ we have, since $\psi(e^{-i\omega}) = \psi(1) = 1$:

$$S_x(\omega) = \frac{\sigma_\varepsilon^2}{\pi|1 - e^{-i\omega}|^2}, \omega \neq 0$$

which shows that the pseudo-spectrum, defined as in theorem 1, is just proportional to the inverse of the Fourier transform of the differencing operator $(1 - L)$ where L is the lag operator. However, if we apply the first difference filter to $\{x_t\}$ the spectrum of $(1 - L)x_t = \varepsilon_t$ is given by $S_\varepsilon(\omega) = \sigma_\varepsilon^2/2\pi$. To perfectly maintain the relation $S_\varepsilon(\omega) = |1 - e^{-i\omega}|^2 S_x(\omega)$ as in the stationary case we would need to define the pseudo-spectrum of x_t as:

$$S_x(\omega) = \frac{\sigma_\varepsilon^2}{2\pi|1 - e^{-i\omega}|^2}, \omega \neq 0$$

which seems a neutral normalization of the (non-integrable) power distribution of x_t . In this case the first difference filter maintains the usual interpretation, summarised by the function $|1 - e^{-i\omega}|^2$. It attenuates low frequencies and amplifies high frequencies, thus producing a "noisier" output series. Now fix $x_0 = 0$. The periodogram of x_t can be written as follows:

$$I_{x,T}(\omega_j) = T^{-1} \left| \sum_{t=1}^T x_t e^{it\omega_j} \right|^2 = \sum_{|k| < T} T^{-1} \sum_{t=1}^{T-|k|} x_t x_{t+|k|} e^{-ik\omega_j}$$

Next, fix any frequency $\omega \in]0, \pi]$. Theorem 1 shows that:

$$\begin{aligned}
E[I_{T,x}(\omega)] &= \sum_{|k|<T} T^{-1} \cos[g(T,\omega)k] \sum_{t=1}^{T-|k|} t = \\
&= \frac{1}{2} \sum_{|k|<T} T^{-1} \cos[g(T,\omega)k] (T-|k|)(T-|k|+1) \rightarrow \frac{\sigma_\varepsilon^2}{\pi|1-e^{-i\omega}|^2} \text{ as } T \rightarrow \infty
\end{aligned}$$

■

Remark 1. Except for $\omega = 0$, the convergence result of theorem 1 does not depend on any initial condition for x_0 . Even when $\omega = 0$ the result seems to be neutral, given the fact that $x_t - x_{t-1}$ is indistinguishable from $(x_t - x_0) - (x_{t-1} - x_0)$. The condition $\sum_{j=-\infty}^{\infty} |\psi_j| |j|^{\frac{1}{2}} < \infty$ is almost always used in a unit-root context but can be relaxed. It is easy to check that the proof works with $\sum_{j=-\infty}^{\infty} |\psi_j| |j|^\alpha < \infty$ for some (small) $\alpha > 0$.

■

Remark 2. A different normalisation is needed for convergence if the order of integration is greater than 1. Consider the simplest case $(1-L)^2 x_t = \varepsilon_t, \forall t$ where $\{\varepsilon_t\}$ is a white noise sequence. Performing the same calculations as in Example 1 we obtain:

$$E[I_{T,x}(\omega)] = \sum_{|k|<T} T^{-1} \cos[g(T,\omega)k] \sum_{t=1}^{T-|k|} t(|k|+t)$$

which diverges since $\sum_{t=1}^{T-|k|} t(|k|+t)$ is a polynomial of order 3 in T . We shall not pursue any frequency domain characterisation in this case.

■

Theorem 1 is an extension of the continuous-time results in Solo (1992). Also, it sharpens the result of theorem 4 in Crato (1996) which gives an upper bound greater than 0 to the limit of (7). In a fractional integration context including unit roots, Hurvich and Ray (1995) have studied the behaviour of the expectation of the periodogram at Fourier frequencies close to the origin, obtaining also a time-invariance result. Specifically, theorem 1 in Hurvich and Ray (1995) shows the following, for a unit-root process:

$$E\left[\frac{1}{2\pi} I_{T,x}(\omega_j) / S_x(\omega_j)\right] \rightarrow 2 \text{ as } T \rightarrow \infty, \omega_j = 2\pi j/T \tag{10}$$

where $S_x(\omega_j)$ is defined as in (12) (see section 3 below), a definition also followed by Velasco (1999). It should be noted that j is held fixed, whereas our result is valid for any fixed $\omega \neq 0$. It is easy to reconcile the two results. Heuristically, once T grows, ω_j approaches 0 and hence $|\psi(e^{-i\omega_j})|^2$ approaches $|\psi(1)|^2$. Therefore $\frac{1}{2\pi}I_{T,x}(\omega_j)$ approaches $2S_x(\omega_j)$, with $S_x(\omega_j)$ defined as in (12). In the stationary case the limit in (10) is just 1.

3 Interpreting filtered integrated time series

If we apply to the stationary sequence $\{x_t\}$ a time-invariant linear filter $h(L) = \sum_{j=-\infty}^{\infty} h_j L^j$, such that $\sum_{j=-\infty}^{\infty} |h_j| < \infty$ we obtain a filtered sequence $y_t = \sum_{k=-\infty}^{\infty} h_k x_{t-k}$. It is easy to verify that the spectrum of $\{y_t\}$ is given by:

$$S_y(\omega) = |h(e^{-i\omega})|^2 S_x(\omega) \quad (11)$$

where $S_x(\omega)$ is the spectrum of $\{x_t\}$ and $h(e^{-i\omega})$ is the transfer function of the filter. Can we extend the relation in (11) to integrated time series? This question is crucial when we want to interpret the effects of applying commonly used moving averages or simply the first difference filter to integrated time series. Common practice is first to define the spectrum of an integrated process as the limit of the spectrum of a stationary process when the smallest autoregressive roots converge to 1 (e.g., Harvey 1993; Den Haan and Sumner 2004; Young, Pedregal and Tych, 1999). For a general ARIMA process the spectrum is defined as:

$$S_x(\omega) = \frac{\sigma_\varepsilon^2}{2\pi} \frac{|\phi^{-1}(e^{-i\omega})|^2 |\theta(e^{-i\omega})|^2}{|1 - e^{-i\omega}|^{2s}} = \frac{\sigma_\varepsilon^2}{2\pi} \frac{|\psi(e^{-i\omega})|^2}{|1 - e^{-i\omega}|^{2s}}, \omega \neq 0 \quad (12)$$

where x_t satisfies:

$$\phi(L)(1 - L)^s x_t = \theta(L)\varepsilon_t, \quad \forall t$$

σ_ε^2 is the variance of the white-noise innovations ε_t , we assume the roots of $\phi(L)$ lie outside the unit circle and are different from those of $\theta(L)$, $\psi(L) = \phi(L)^{-1}\theta(L)$ and $s \geq 0$ the order of integration of the series. This limit is a time-invariant function at all frequencies except at those associated with autoregressive roots with unit modulus¹. An extension of the relation in (11) holds given the definition in (12), particularly when the filter renders the series stationary. It

¹Since we assumed the roots of $\phi(L)$ lie outside the unit circle, we are only considering the existence of a pole at zero frequency. This assumption can straightforwardly be relaxed in order to include singularities at frequencies other than zero, e.g., due to non-stationary seasonal components.

is assumed, without resorting to results such as that in theorem 1, that this function represents indeed a distribution of variance.

Bujosa, Bujosa and García-Ferrer (2002) provide a rigorous justification to the definition in (12), generalising the classical spectral analysis by developing an extended Fourier transform to the field of fractions of polynomials. A pseudo-autocovariance generating function is defined and the corresponding extended Fourier transform is defined as the (pseudo-) spectrum of the integrated series, which leads to a functional form exactly as in (12). No representation theorem is provided but it is argued, again without stating a result such as that in theorem 1, that the usual interpretation of the spectrum as a decomposition of variance holds. Were the functions in theorem 1 (which only deals with one unit root) and in (12) the same for $\omega \neq 0$, one could state that defining the spectrum of an integrated series as the limit of the expected periodogram was a coherent extension of the stationary case inversion relation in (2). But the alert reader has noticed that the functional form in theorem 1 is slightly different than that in (12) due to the term $|\psi(1)|^2$ in the numerator. This is definitely a nuisance when the process is not a pure random walk, for which a straightforward normalisation (as in Example 1) preserves the power distribution and leads to the maintenance of the relation in (11). In any case, the differences in the interpretation would not be dramatic given the fact that the inverse of $|1 - e^{-i\omega}|^2$ dominates the behaviour of both functions at frequencies close to the pole located at zero frequency. It seems that a special mean correction would be needed to eliminate $|\psi(1)|^2$ from the numerator, as Solo (1992) argues in the context of the dependence of his theorem 1 on initial conditions. We shall not pursue it here.

4 Conclusions

We have shown that the limit of the expected periodogram is a time invariant function when the time series process contains one unit root, exactly as in the stationary case. One could be tempted to define this function as the spectrum (or power distribution) of the integrated process, since in the stationary case this limit is the distribution of power revealed by the spectral representation theorem. However, this function differs slightly from the commonly defined (pseudo-) spectrum of an integrated time series, which has recently been given a rigorous interpretation by Bujosa, Bujosa and García-Ferrer (2002). Defining the spectrum of an integrated series as in theorem 1 would in general distort the interpretation given to the transfer function of filters applied to such series.

References

- [1] Bujosa, A., Bujosa, M. and García-Ferrer (2002). A Note on Pseudo-spectra and Pseudo-Covariance Generating Functions of ARMA processes. mimeo
- [2] Crato, N. (1996). Some results on the spectral analysis of nonstationary time series. *Portugaliae Mathematica*, 53(2):179:186
- [3] Den Haan, W.J., and Sumner, S.W. (2004). The Comovements between Real Activity and Prices in the G7, *European Economic Review*, 48:1333-1347
- [4] Harvey, A.C. (1993). *Forecasting, structural time series models and the Kalman filter*. Cambridge University Press
- [5] Hurvich, C.M. and Ray, B.K., (1995). Estimation of the memory parameter for nonstationary or noninvertible fractionally integrated processes. *Journal of Time Series Analysis*, 16:17-42.
- [6] Solo, V.(1992). Intrinsic random functions and the paradox of $1/f$ noise. *SIAM Journal of Applied Mathematics*, 52(1):270-291
- [7] Velasco, C., (1999). Non-stationary log-periodogram regression. *Journal of Econometrics*, 91:325-371.
- [8] Young, P. C., Pedregal, D. and Tych, W. (1999). Dynamic Harmonic Regression. *Journal of Forecasting*, 18:369-394.

WORKING PAPERS

2000

- 1/00 UNEMPLOYMENT DURATION: COMPETING AND DEFECTIVE RISKS
— *John T. Addison, Pedro Portugal*
- 2/00 THE ESTIMATION OF RISK PREMIUM IMPLICIT IN OIL PRICES
— *Jorge Barros Luís*
- 3/00 EVALUATING CORE INFLATION INDICATORS
— *Carlos Robalo Marques, Pedro Duarte Neves, Luís Morais Sarmiento*
- 4/00 LABOR MARKETS AND KALEIDOSCOPIIC COMPARATIVE ADVANTAGE
— *Daniel A. Traça*
- 5/00 WHY SHOULD CENTRAL BANKS AVOID THE USE OF THE UNDERLYING INFLATION INDICATOR?
— *Carlos Robalo Marques, Pedro Duarte Neves, Afonso Gonçalves da Silva*
- 6/00 USING THE ASYMMETRIC TRIMMED MEAN AS A CORE INFLATION INDICATOR
— *Carlos Robalo Marques, João Machado Mota*

2001

- 1/01 THE SURVIVAL OF NEW DOMESTIC AND FOREIGN OWNED FIRMS
— *José Mata, Pedro Portugal*
- 2/01 GAPS AND TRIANGLES
— *Bernardino Adão, Isabel Correia, Pedro Teles*
- 3/01 A NEW REPRESENTATION FOR THE FOREIGN CURRENCY RISK PREMIUM
— *Bernardino Adão, Fátima Silva*
- 4/01 ENTRY MISTAKES WITH STRATEGIC PRICING
— *Bernardino Adão*
- 5/01 FINANCING IN THE EUROSISTEM: FIXED VERSUS VARIABLE RATE TENDERS
— *Margarida Catalão-Lopes*
- 6/01 AGGREGATION, PERSISTENCE AND VOLATILITY IN A MACROMODEL
— *Karim Abadir, Gabriel Talmain*
- 7/01 SOME FACTS ABOUT THE CYCLICAL CONVERGENCE IN THE EURO ZONE
— *Frederico Belo*
- 8/01 TENURE, BUSINESS CYCLE AND THE WAGE-SETTING PROCESS
— *Leandro Arozamena, Mário Centeno*
- 9/01 USING THE FIRST PRINCIPAL COMPONENT AS A CORE INFLATION INDICATOR
— *José Ferreira Machado, Carlos Robalo Marques, Pedro Duarte Neves, Afonso Gonçalves da Silva*
- 10/01 IDENTIFICATION WITH AVERAGED DATA AND IMPLICATIONS FOR HEDONIC REGRESSION STUDIES
— *José A.F. Machado, João M.C. Santos Silva*

2002

- 1/02 QUANTILE REGRESSION ANALYSIS OF TRANSITION DATA
— José A.F. Machado, Pedro Portugal
- 2/02 SHOULD WE DISTINGUISH BETWEEN STATIC AND DYNAMIC LONG RUN EQUILIBRIUM IN ERROR CORRECTION MODELS?
— Susana Botas, Carlos Robalo Marques
- 3/02 MODELLING TAYLOR RULE UNCERTAINTY
— Fernando Martins, José A. F. Machado, Paulo Soares Esteves
- 4/02 PATTERNS OF ENTRY, POST-ENTRY GROWTH AND SURVIVAL: A COMPARISON BETWEEN DOMESTIC AND FOREIGN OWNED FIRMS
— José Mata, Pedro Portugal
- 5/02 BUSINESS CYCLES: CYCLICAL COMOVEMENT WITHIN THE EUROPEAN UNION IN THE PERIOD 1960-1999. A FREQUENCY DOMAIN APPROACH
— João Valle e Azevedo
- 6/02 AN “ART”, NOT A “SCIENCE”? CENTRAL BANK MANAGEMENT IN PORTUGAL UNDER THE GOLD STANDARD, 1854 -1891
— Jaime Reis
- 7/02 MERGE OR CONCENTRATE? SOME INSIGHTS FOR ANTITRUST POLICY
— Margarida Catalão-Lopes
- 8/02 DISENTANGLING THE MINIMUM WAGE PUZZLE: ANALYSIS OF WORKER ACCESSIONS AND SEPARATIONS FROM A LONGITUDINAL MATCHED EMPLOYER-EMPLOYEE DATA SET
— Pedro Portugal, Ana Rute Cardoso
- 9/02 THE MATCH QUALITY GAINS FROM UNEMPLOYMENT INSURANCE
— Mário Centeno
- 10/02 HEDONIC PRICES INDEXES FOR NEW PASSENGER CARS IN PORTUGAL (1997-2001)
— Hugo J. Reis, J.M.C. Santos Silva
- 11/02 THE ANALYSIS OF SEASONAL RETURN ANOMALIES IN THE PORTUGUESE STOCK MARKET
— Miguel Balbina, Nuno C. Martins
- 12/02 DOES MONEY GRANGER CAUSE INFLATION IN THE EURO AREA?
— Carlos Robalo Marques, Joaquim Pina
- 13/02 INSTITUTIONS AND ECONOMIC DEVELOPMENT: HOW STRONG IS THE RELATION?
— Tiago V.de V. Cavalcanti, Álvaro A. Novo

2003

- 1/03 FOUNDING CONDITIONS AND THE SURVIVAL OF NEW FIRMS
— P.A. Geroski, José Mata, Pedro Portugal
- 2/03 THE TIMING AND PROBABILITY OF FDI: AN APPLICATION TO THE UNITED STATES MULTINATIONAL ENTERPRISES
— José Brandão de Brito, Felipa de Mello Sampayo
- 3/03 OPTIMAL FISCAL AND MONETARY POLICY: EQUIVALENCE RESULTS
— Isabel Correia, Juan Pablo Nicolini, Pedro Teles

- 4/03** FORECASTING EURO AREA AGGREGATES WITH BAYESIAN VAR AND VECM MODELS
— *Ricardo Mourinho Félix, Luís C. Nunes*
- 5/03** CONTAGIOUS CURRENCY CRISES: A SPATIAL PROBIT APPROACH
— *Álvaro Novo*
- 6/03** THE DISTRIBUTION OF LIQUIDITY IN A MONETARY UNION WITH DIFFERENT PORTFOLIO RIGIDITIES
— *Nuno Alves*
- 7/03** COINCIDENT AND LEADING INDICATORS FOR THE EURO AREA: A FREQUENCY BAND APPROACH
— *António Rua, Luís C. Nunes*
- 8/03** WHY DO FIRMS USE FIXED-TERM CONTRACTS?
— *José Varejão, Pedro Portugal*
- 9/03** NONLINEARITIES OVER THE BUSINESS CYCLE: AN APPLICATION OF THE SMOOTH TRANSITION AUTOREGRESSIVE MODEL TO CHARACTERIZE GDP DYNAMICS FOR THE EURO-AREA AND PORTUGAL
— *Francisco Craveiro Dias*
- 10/03** WAGES AND THE RISK OF DISPLACEMENT
— *Anabela Carneiro, Pedro Portugal*
- 11/03** SIX WAYS TO LEAVE UNEMPLOYMENT
— *Pedro Portugal, John T. Addison*
- 12/03** EMPLOYMENT DYNAMICS AND THE STRUCTURE OF LABOR ADJUSTMENT COSTS
— *José Varejão, Pedro Portugal*
- 13/03** THE MONETARY TRANSMISSION MECHANISM: IS IT RELEVANT FOR POLICY?
— *Bernardino Adão, Isabel Correia, Pedro Teles*
- 14/03** THE IMPACT OF INTEREST-RATE SUBSIDIES ON LONG-TERM HOUSEHOLD DEBT: EVIDENCE FROM A LARGE PROGRAM
— *Nuno C. Martins, Ernesto Villanueva*
- 15/03** THE CAREERS OF TOP MANAGERS AND FIRM OPENNESS: INTERNAL VERSUS EXTERNAL LABOUR MARKETS
— *Francisco Lima, Mário Centeno*
- 16/03** TRACKING GROWTH AND THE BUSINESS CYCLE: A STOCHASTIC COMMON CYCLE MODEL FOR THE EURO AREA
— *João Valle e Azevedo, Siem Jan Koopman, António Rua*
- 17/03** CORRUPTION, CREDIT MARKET IMPERFECTIONS, AND ECONOMIC DEVELOPMENT
— *António R. Antunes, Tiago V. Cavalcanti*
- 18/03** BARGAINED WAGES, WAGE DRIFT AND THE DESIGN OF THE WAGE SETTING SYSTEM
— *Ana Rute Cardoso, Pedro Portugal*
- 19/03** UNCERTAINTY AND RISK ANALYSIS OF MACROECONOMIC FORECASTS: FAN CHARTS REVISITED
— *Álvaro Novo, Maximiano Pinheiro*
- 2004**
- 1/04** HOW DOES THE UNEMPLOYMENT INSURANCE SYSTEM SHAPE THE TIME PROFILE OF JOBLESS DURATION?
— *John T. Addison, Pedro Portugal*

- 2/04** REAL EXCHANGE RATE AND HUMAN CAPITAL IN THE EMPIRICS OF ECONOMIC GROWTH
— *Delfim Gomes Neto*
- 3/04** ON THE USE OF THE FIRST PRINCIPAL COMPONENT AS A CORE INFLATION INDICATOR
— *José Ramos Maria*
- 4/04** OIL PRICES ASSUMPTIONS IN MACROECONOMIC FORECASTS: SHOULD WE FOLLOW FUTURES MARKET EXPECTATIONS?
— *Carlos Coimbra, Paulo Soares Esteves*
- 5/04** STYLISED FEATURES OF PRICE SETTING BEHAVIOUR IN PORTUGAL: 1992-2001
— *Mónica Dias, Daniel Dias, Pedro D. Neves*
- 6/04** A FLEXIBLE VIEW ON PRICES
— *Nuno Alves*
- 7/04** ON THE FISHER-KONIECZNY INDEX OF PRICE CHANGES SYNCHRONIZATION
— *D.A. Dias, C. Robalo Marques, P.D. Neves, J.M.C. Santos Silva*
- 8/04** INFLATION PERSISTENCE: FACTS OR ARTEFACTS?
— *Carlos Robalo Marques*
- 9/04** WORKERS' FLOWS AND REAL WAGE CYCLICALITY
— *Anabela Carneiro, Pedro Portugal*
- 10/04** MATCHING WORKERS TO JOBS IN THE FAST LANE: THE OPERATION OF FIXED-TERM CONTRACTS
— *José Varejão, Pedro Portugal*
- 11/04** THE LOCATIONAL DETERMINANTS OF THE U.S. MULTINATIONALS ACTIVITIES
— *José Brandão de Brito, Felipa Mello Sampayo*
- 12/04** KEY ELASTICITIES IN JOB SEARCH THEORY: INTERNATIONAL EVIDENCE
— *John T. Addison, Mário Centeno, Pedro Portugal*
- 13/04** RESERVATION WAGES, SEARCH DURATION AND ACCEPTED WAGES IN EUROPE
— *John T. Addison, Mário Centeno, Pedro Portugal*
- 14/04** THE MONETARY TRANSMISSION IN THE US AND THE EURO AREA: COMMON FEATURES AND COMMON FRICTIONS
— *Nuno Alves*
- 15/04** NOMINAL WAGE INERTIA IN GENERAL EQUILIBRIUM MODELS
— *Nuno Alves*
- 16/04** MONETARY POLICY IN A CURRENCY UNION WITH NATIONAL PRICE ASYMMETRIES
— *Sandra Gomes*
- 17/04** NEOCLASSICAL INVESTMENT WITH MORAL HAZARD
— *João Ejarque*
- 18/04** MONETARY POLICY WITH STATE CONTINGENT INTEREST RATES
— *Bernardino Adão, Isabel Correia, Pedro Teles*
- 19/04** MONETARY POLICY WITH SINGLE INSTRUMENT FEEDBACK RULES
— *Bernardino Adão, Isabel Correia, Pedro Teles*
- 20/04** ACCOUNTING FOR THE HIDDEN ECONOMY: BARRIERS TO LEGALITY AND LEGAL FAILURES
— *António R. Antunes, Tiago V. Cavalcanti*

2005

- 1/05** SEAM: A SMALL-SCALE EURO AREA MODEL WITH FORWARD-LOOKING ELEMENTS
— *José Brandão de Brito, Rita Duarte*
- 2/05** FORECASTING INFLATION THROUGH A BOTTOM-UP APPROACH: THE PORTUGUESE CASE
— *Cláudia Duarte, António Rua*
- 3/05** USING MEAN REVERSION AS A MEASURE OF PERSISTENCE
— *Daniel Dias, Carlos Robalo Marques*
- 4/05** HOUSEHOLD WEALTH IN PORTUGAL: 1980-2004
— *Fátima Cardoso, Vanda Geraldés da Cunha*
- 5/05** ANALYSIS OF DELINQUENT FIRMS USING MULTI-STATE TRANSITIONS
— *António Antunes*
- 6/05** PRICE SETTING IN THE AREA: SOME STYLIZED FACTS FROM INDIVIDUAL CONSUMER PRICE DATA
— *Emmanuel Dhyne, Luis J. Álvarez, Hervé Le Bihan, Giovanni Veronese, Daniel Dias, Johannes Hoffmann, Nicole Jonker, Patrick Lünnemann, Fabio Ruml, Jouko Vilmunen*
- 7/05** INTERMEDIATION COSTS, INVESTOR PROTECTION AND ECONOMIC DEVELOPMENT
— *António Antunes, Tiago Cavalcanti, Anne Villamil*
- 8/05** TIME OR STATE DEPENDENT PRICE SETTING RULES? EVIDENCE FROM PORTUGUESE MICRO DATA
— *Daniel Dias, Carlos Robalo Marques, João Santos Silva*
- 9/05** BUSINESS CYCLE AT A SECTORAL LEVEL: THE PORTUGUESE CASE
— *Hugo Reis*
- 10/05** THE PRICING BEHAVIOUR OF FIRMS IN THE EURO AREA: NEW SURVEY EVIDENCE
— *S. Fabiani, M. Druant, I. Hernando, C. Kwapił, B. Landau, C. Loupias, F. Martins, T. Mathä, R. Sabbatini, H. Stahl, A. Stokman*
- 11/05** CONSUMPTION TAXES AND REDISTRIBUTION
— *Isabel Correia*
- 12/05** UNIQUE EQUILIBRIUM WITH SINGLE MONETARY INSTRUMENT RULES
— *Bernardino Adão, Isabel Correia, Pedro Teles*
- 13/05** A MACROECONOMIC STRUCTURAL MODEL FOR THE PORTUGUESE ECONOMY
— *Ricardo Mourinho Félix*
- 14/05** THE EFFECTS OF A GOVERNMENT EXPENDITURES SHOCK
— *Bernardino Adão, José Brandão de Brito*
- 15/05** MARKET INTEGRATION IN THE GOLDEN PERIPHERY – THE LISBON/LONDON EXCHANGE, 1854-1891
— *Rui Pedro Esteves, Jaime Reis, Fabiano Ferramosca*

2006

- 1/06** THE EFFECTS OF A TECHNOLOGY SHOCK IN THE EURO AREA
— *Nuno Alves, José Brandão de Brito, Sandra Gomes, João Sousa*
- 2/02** THE TRANSMISSION OF MONETARY AND TECHNOLOGY SHOCKS IN THE EURO AREA
— *Nuno Alves, José Brandão de Brito, Sandra Gomes, João Sousa*

- 3/06** MEASURING THE IMPORTANCE OF THE UNIFORM NONSYNCHRONIZATION HYPOTHESIS
— *Daniel Dias, Carlos Robalo Marques, João Santos Silva*
- 4/06** THE PRICE SETTING BEHAVIOUR OF PORTUGUESE FIRMS EVIDENCE FROM SURVEY DATA
— *Fernando Martins*
- 5/06** STICKY PRICES IN THE EURO AREA: A SUMMARY OF NEW MICRO EVIDENCE
— *L. J. Álvarez, E. Dhyne, M. Hoeberichts, C. Kwapil, H. Le Bihan, P. Lünemann, F. Martins, R. Sabbatini, H. Stahl, P. Vermeulen and J. Vilmunen*
- 6/06** NOMINAL DEBT AS A BURDEN ON MONETARY POLICY
— *Javier Díaz-Giménez, Giorgia Giovannetti, Ramon Marimon, Pedro Teles*
- 7/06** A DISAGGREGATED FRAMEWORK FOR THE ANALYSIS OF STRUCTURAL DEVELOPMENTS IN PUBLIC FINANCES
— *Jana Kremer, Cláudia Rodrigues Braz, Teunis Brosens, Geert Langenus, Sandro Momigliano, Mikko Spolander*
- 8/06** IDENTIFYING ASSET PRICE BOOMS AND BUSTS WITH QUANTILE REGRESSIONS
— *José A. F. Machado, João Sousa*
- 9/06** EXCESS BURDEN AND THE COST OF INEFFICIENCY IN PUBLIC SERVICES PROVISION
— *António Afonso, Vítor Gaspar*
- 10/06** MARKET POWER, DISMISSAL THREAT AND RENT SHARING: THE ROLE OF INSIDER AND OUTSIDER FORCES IN WAGE BARGAINING
— *Anabela Carneiro, Pedro Portugal*
- 11/06** MEASURING EXPORT COMPETITIVENESS: REVISITING THE EFFECTIVE EXCHANGE RATE WEIGHTS FOR THE EURO AREA COUNTRIES
— *Paulo Soares Esteves, Carolina Reis*
- 12/06** THE IMPACT OF UNEMPLOYMENT INSURANCE GENEROSITY ON MATCH QUALITY DISTRIBUTION
— *Mário Centeno, Alvaro A. Novo*
- 13/06** U.S. UNEMPLOYMENT DURATION: HAS LONG BECOME LONGER OR SHORT BECOME SHORTER?
— *José A.F. Machado, Pedro Portugal e Juliana Guimarães*
- 14/06** EARNINGS LOSSES OF DISPLACED WORKERS: EVIDENCE FROM A MATCHED EMPLOYER-EMPLOYEE DATA SET
— *Anabela Carneiro, Pedro Portugal*
- 15/06** COMPUTING GENERAL EQUILIBRIUM MODELS WITH OCCUPATIONAL CHOICE AND FINANCIAL FRICTIONS
— *António Antunes, Tiago Cavalcanti, Anne Villamil*
- 16/06** ON THE RELEVANCE OF EXCHANGE RATE REGIMES FOR STABILIZATION POLICY
— *Bernardino Adao, Isabel Correia, Pedro Teles*
- 17/06** AN INPUT-OUTPUT ANALYSIS: LINKAGES VS LEAKAGES
— *Hugo Reis, António Rua*
- 2007**
- 1/07** RELATIVE EXPORT STRUCTURES AND VERTICAL SPECIALIZATION: A SIMPLE CROSS-COUNTRY INDEX
— *João Amador, Sónia Cabral, José Ramos Maria*

- 2/07** THE FORWARD PREMIUM OF EURO INTEREST RATES
— *Sónia Costa, Ana Beatriz Galvão*
- 3/07** ADJUSTING TO THE EURO
— *Gabriel Fagan, Vítor Gaspar*
- 4/07** SPATIAL AND TEMPORAL AGGREGATION IN THE ESTIMATION OF LABOR DEMAND FUNCTIONS
— *José Varejão, Pedro Portugal*
- 5/07** PRICE SETTING IN THE EURO AREA: SOME STYLISED FACTS FROM INDIVIDUAL PRODUCER PRICE DATA
— *Philip Vermeulen, Daniel Dias, Maarten Dossche, Erwan Gautier, Ignacio Hernando, Roberto Sabbatini, Harald Stahl*
- 6/07** A STOCHASTIC FRONTIER ANALYSIS OF SECONDARY EDUCATION OUTPUT IN PORTUGAL
— *Manuel Coutinho Pereira, Sara Moreira*
- 7/07** CREDIT RISK DRIVERS: EVALUATING THE CONTRIBUTION OF FIRM LEVEL INFORMATION AND OF MACROECONOMIC DYNAMICS
— *Diana Bonfim*
- 8/07** CHARACTERISTICS OF THE PORTUGUESE ECONOMIC GROWTH: WHAT HAS BEEN MISSING?
— *João Amador, Carlos Coimbra*
- 9/07** TOTAL FACTOR PRODUCTIVITY GROWTH IN THE G7 COUNTRIES: DIFFERENT OR ALIKE?
— *João Amador, Carlos Coimbra*
- 10/07** IDENTIFYING UNEMPLOYMENT INSURANCE INCOME EFFECTS WITH A QUASI-NATURAL EXPERIMENT
— *Mário Centeno, Alvaro A. Novo*
- 11/07** HOW DO DIFFERENT ENTITLEMENTS TO UNEMPLOYMENT BENEFITS AFFECT THE TRANSITIONS FROM UNEMPLOYMENT INTO EMPLOYMENT
— *John T. Addison, Pedro Portugal*
- 12/07** INTERPRETATION OF THE EFFECTS OF FILTERING INTEGRATED TIME SERIES
— *João Valle e Azevedo*
- 13/07** EXACT LIMIT OF THE EXPECTED PERIODOGRAM IN THE UNIT-ROOT CASE
— *João Valle e Azevedo*